

SUCCESS IS SOMETHING TO SNEEZE AT

Influenza Mortality in Cities that Participate in the Super Bowl

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ABSTRACT

Using county-level Vital Statistics of the United States data from 1974 to 2009, we employ a differences-in-differences framework comparing influenza mortality rates in Bowl-participating counties to nonparticipants. We estimate having a local team in the Super Bowl caused an 18 percent increase in influenza deaths for the population over age 65. Results are most pronounced in years when the dominant influenza strain is more virulent, or when the Super Bowl occurs closer to the peak of influenza season. We find no impacts on influenza mortality in hosting cities. Our findings suggest mitigating transmission at gatherings related to large spectator events could have substantial returns for public health.

KEYWORDS: influenza, externalities, gatherings, Super Bowl

JEL CLASSIFICATION: D62, I12, L83

I. Introduction

Governments dedicate substantial public health resources to containment of influenza pandemics. For example, the federal government spent \$3 billion in direct response to the 2009 H1N1 pandemic (Centers for Disease Control and Prevention, unpublished data). Uncontained epidemics can be costly in terms of loss of life: the 1918–19 and 1957–59 pandemics caused an estimated 500,000 and 86,000 excess deaths, respectively, in the United States (Crosby 2003). The Centers for Disease Control and Prevention (CDC) estimates a pandemic of similar magnitude to that in 1957 would cost the United States \$71 billion today, even before including disruptions to commerce and society (Meltzer, Cox, and Fukuda 1999). A more recent estimate puts the cost of the 2009 H1N1 pandemic at \$65 billion, including all societal costs (Lee et al. 2010). However, these total cost calculations may be underestimates as they do not include the long-term costs of *in utero* exposure including increased rates of disability (Almond 2006), chronic conditions (Lin and Liu 2014), and cognitive development (Kelly 2011). Investments made *ex ante* that minimize risk may be efficient given such large *ex post* containment costs. Vaccination represents the most effective *preventive* strategy, but is often unavailable in the short term because

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of production constraints (Hinman et al. 2006; Sloan et al. 2004). Changes in mixing patterns (the degree and manner of physical interaction among individuals) may provide an alternate short-term shift in transmission rates. We use annual and geographic variation in local sports team success as a proxy for changes in mixing rates through increased fan travel or local gatherings. We find increased mixing, as proxied by team success, increases local influenza mortality, suggesting small-scale mixing plays an important role.

Influenza is an infectious disease that spreads by airborne droplets with an approximate travel radius of 6 feet, making close human contact an important infection vector (Centers for Disease Control and Prevention 2012). Strategies for transmission reduction may include altering human mixing rates and limited contact with “vulnerable” populations such as the elderly or very young, but there is limited empirical evidence changes in mixing and contact have an economically meaningful impact on influenza incidence. Recent studies rely on plausibly exogenous events as natural experiments that alter contact or travel rates. For example, Brownstein, Wolfe, and Mandl (2006) examine how the decline in air travel after September 11, 2001, affected the seasonal influenza peak, and found a 27 percent reduction in air travel postponed the peak by 13 days.¹ Cauchemez et al. (2008) examine the impact of holiday school closures in France on influenza transmission in children. Research detected heightened influenza infection rates at the 2002 Winter Olympics (Gundlapalli et al. 2006), the 2006 World Cup (Schenkel et al. 2006), music festivals (Botelho-Nevers et al. 2010; Gutiérrez et al. 2009), and the yearly Hajj pilgrimage to Mecca (Balkhy et al. 2004). Such studies are informative, but focus only on a single instance of each event, thus lacking explicit control groups.

In contrast, we consider influenza transmission in the context of a repeated event with explicit treatment and control groups. We examine how local sports team success, here participation in the championship game of the National Football League (NFL), impacts influenza mortality in the United States. The Super Bowl takes place every year in either January or February, traditionally high influenza periods, providing more than 30 years of similar events as a basis for our study. We use team success as a proxy for shifts in socializing behavior during periods of traditionally higher influenza transmission. Using mortality data from restricted Multiple Causes of Death data covering 1974–2009, we calculate percentage changes in reported influenza deaths, with a focus on mortality for those over the age of 65, the most vulnerable population historically (Thompson et al. 2003). Though populations over 65 may not change behavior in response to team success, their chances of contact with an infected individual increases as the infection rate increases in the population overall.² Our core estimation strategy employs a differences-in-differences

1 This result may be particular to 2001 and has not been borne out by correlations in 30 other years of data (Viboud et al. 2006).

2 Vaccinating health-care workers is a demonstrated strategy for reducing influenza mortality in the elderly (Potter et al. 1997), which shows a potential transmission path from more sports-social age groups to potentially less enthusiastic older populations. Further, while daily contacts are generally with individuals of similar ages, this assortative mixing within age groups is least pronounced for those age 65 and older (Mosson et al. 2008).

FIGURE 1. Frequency of participating in the Super Bowl, 1974–2009

Notes: Size of the dot indicates relative number of times the city participated in the Super Bowl between 1974 and 2009. Larger dots indicate a greater participation frequency.

framework to compare areas with successful football teams in a given year, as measured by qualifying for the Super Bowl, to areas that also house football teams but did not successfully reach the Super Bowl in that year.

There are several advantages to using Super Bowl attendance as a proxy for increased travel and gatherings. First, the championship game itself is a very large gathering; in recent years, attendance has crested 100,000. Second, the postseason often takes place during the height of influenza season (influenza morbidity and mortality generally peaks between December and March in the United States (Doshi 2008)), making related travel and gatherings more relevant for our disease of interest. Third, cities with NFL teams that did not qualify for postseason play make an ideal control group, as they are relatively similar in size, wealth, and influenza mortality patterns to those that attend. This is particularly true in professional football, where revenue sharing and salary cap policies reduce the formation of legacy teams that annually attend the Super Bowl; Figure 1 shows that for the 36 championship games between 1974 and 2009, 25 of the 31 NFL cities had a team qualify at least once. Fourth, the location of postseason play changes from year to year. For playoff games, relative team records determine the hosting city. Super Bowl games occur in predetermined cities, some of which are frequent hosts—Figure 2 shows that all 36 seasons we consider culminate in the same 11 host cities.

We hypothesize three plausible mechanisms through which a local team qualifying for postseason play could impact influenza transmission. First, an increase in large gatherings increases the frequency of human contact and probability of transmission. Here, postseason play can increase gatherings both in the host area (through the playoff or championship game itself) and in the areas with successful teams (through increased attendance at home sports parties, sports bars, etc.). Second, postseason play alters travel patterns and increases the contact or mixing rate between possible susceptible and infected groups via fan mobility, increasing opportunities for exchange as well as co-infection. Third,

FIGURE 2. Frequency of hosting the Super Bowl, 1974–2009

Notes: Size of the dot indicates relative number of times the city hosted the Super Bowl between 1974 and 2009. Larger dots indicate a greater hosting frequency.

postseason play may affect local economies in areas with participating teams, through increased tourism from outside or increased local expenditures. Changes in local income or employment rates may in turn influence behavior and health expenditures. Hosting a championship game moves workers from other sectors of the economy into the amusements and recreation sector (Coates and Humphreys 2003), altering local business patterns and large-scale interaction. Without specific data on travel and gatherings, we are unable to explicitly separate the three potential mechanisms. However, we consider both (1) mortality in cities that host the Super Bowl and (2) mortality in cities that send a team to the Super Bowl. Using annual variation in influenza mortality risk across both sending and hosting cities, we find strong evidence in favor of local gathering transmission with little evidence of travel or economic effects: hosting the Super Bowl does little to change mortality risk in the host metropolitan statistical area (MSA), and sending-team MSA outcomes correlate more with local risk rates than hosting-MSA risk rates.

We contribute a unique perspective to the literature on benefits of both promoting local sports teams as a source of economic development and hosting large-scale sports events as a source of tourism. As of yet, no other studies examine the potential disease costs imposed on local residents from having a successful sports team. We find counties in MSAs with teams participating in championship play see an increase in influenza deaths of approximately 18 percent, with larger effects when the game occurs closer to the peak of influenza season, or when the dominant influenza strain is more lethal. We find nonzero effects for teams in playoff games as well, which suggests some effect comes from general team success and supports the mechanism of increased local gatherings. Interestingly, we find no increases in mortality in the host cities themselves.³ As is expected, we find no

3 Controversy about the true size of the economic impact of being championship game host city has proliferated in the popular press (Matheson and Baade 2006; Bialik 2010; Rische 2012), and increased influenza transmission presents another potential externality of being a host city.

comparable impacts for noncommunicable diseases or for influenza transmission one or two years before or after the successful year.

II. Estimation Strategy

We employ a differences-in-differences model, where the unit of observation is the county-season, to identify the role of team success on local health. In a given season, treated counties are those that reside within a metropolitan statistical area (MSA) housing an NFL team that made it to the Super Bowl. Control counties are those in MSAs housing less successful, non-Bowl NFL teams. Though treatment and control status is common across counties within an MSA, we conduct our analysis at the county level to more precisely control for time-varying local conditions. Specifically, we estimate

$$Mort_{c,t} = \alpha + \beta Attend_{c,t} + \delta X_{c,t} + \lambda_c + \gamma_t + \epsilon_{c,t} \quad (1),$$

where $Mort_{c,t}$ is the influenza mortality rate in county c at season t ; $Attend_{c,t}$ is a dummy variable equal to one when the team in the MSA including county c makes it to the Super Bowl; and $X_{c,t}$ is a collection of county-specific controls at time t . To control for demographic changes that may make a county more susceptible on average to influenza mortality (Quinn et al. 2011), we control for percentage of the population that is nonwhite. While the majority of our regressions explore death rates in the population age 65 and over, we also control directly for the percentage of the population age 65 and older to account for potential changes in contact patterns within counties. To account for fluctuations in influenza transmission due to environmental factors (Soebiyanto, Adimi, and Kiang 2010), which could be potentially correlated with local team success through home field effects or more productive practice sessions, we include flexible time-varying environmental controls. We control for average number of days in bins of 10 degrees Fahrenheit with categories for less than or equal to 30 degrees or greater than 90 degrees, average number of days with humidity in 2 grams per kilogram (grams of water vapor per kilogram of air) bins with a category for greater than 18 grams per kilogram, and average number of days with precipitation in 0.5-inch bins with a category for greater than 1 inch of precipitation.

The coefficients λ and γ are county and season of influenza fixed effects to account for different background levels of disease activity in each place or season. In expanded models, we include county-specific linear or quadratic time trends to allow for correlation between changing health outcomes and the likelihood of sending a team to the Super Bowl over time (though as noted above, profit-sharing among NFL teams reduces correlation in team success across time). The coefficient β shows the reduced-form marginal effect of sending a team to the Super Bowl on influenza mortality (relative to other counties in MSAs that house NFL teams). In models checking for local hosting effects, we classify treatment as being in the same MSA as the host stadium of the Super Bowl. We weight all regressions by population of the indicated age cell, and cluster standard errors at the MSA level to account for geographically common unexplained relationships between local sports team success and influenza mortality.

III. Data

We combine data on death records, population, and weather with dates and locations of Super Bowls between 1974 and 2009.⁴ The number and location of deaths come from restricted use Multiple Cause of Death Files from the National Center for Health Statistics Vital Statistics System.⁵ We generate the denominator for death rates using population data from the Surveillance, Epidemiology, and End Results (SEER) Program from the National Cancer Institute. We use the county-level file with data on population by race and single-year age group. We collapse all estimates to the county by relevant age group indicated in the tables. To match population denominators, we assign deaths by county of residence rather than county of occurrence.⁶

The Multiple Cause of Death File data show cause of death according to International Classification of Disease (ICD) codes that appear on the death certificate. We code deaths as due to influenza if any of the influenza codes appear on the death certificate. This means we incorrectly classify deaths due to influenza but marked as another cause, causing measurement error in our outcome of interest. In effect, we explore the relationship between sending a team to the Super Bowl and *observed* influenza mortality rather than the total number of influenza deaths. Assuming any such error is classical and uncorrelated with unobservables, it will increase standard error levels making statistical inference more difficult. Our controls and fixed effects aim to capture and remove further differences in levels or trends of disease classification that vary across counties or over time in a manner unrelated to the Super Bowl. As our period of interest spans three ICD editions (8, 9, and 10), we follow published comparability classifications (Klebbba and Scott 1980; Anderson et al. 2001) to maintain consistency across systems.⁷ In some estimates we stratify our model by dominant influenza subtype using previously published data for seasons from 1974 to 1999 (Greene, Ionides, and Wilson 2006) and from Influenza Season Summaries provided by the Centers for Disease Control and Prevention for seasons after 1999 (Centers for Disease Control and Prevention 2014).

To control for weather effects, we obtain county-level humidity measurements from the Global Summary of the Day and temperatures and precipitation measures from the Global Historical Climatology Network. The National Oceanic and Atmospheric Administration National Climatic Data Center maintains both data sets.

4 We omit the 1972–73 flu season because of a disproportionately high mortality in California that year, which also served as the hosting state for that year's Super Bowl (Chin, Magoffin, and Lennette 1974). Results including 1972–1973 show unrealistically large effects of being a hosting county that persist in no other year.

5 Starting in 1989 county of residence for small counties and date of death are only available in the restricted use file: tabulations of MSA-level mortality rates (as well as compilations of Super Bowl and playoff hosts, participants, and dates) are available upon request.

6 County of occurrence may correlate with the Super Bowl for reasons other than disease transmission caused by the event. A decedent that traveled to a county associated with a successful NFL team may contribute to the number of influenza deaths in that county, but he may have died regardless.

7 Influenza deaths are codes 470–474 in the Eighth Revision, 487 in the Ninth Revision, and J10–J11 in the Tenth Revision.

TABLE 1. Summary statistics for counties associated with the NFL

	(1) Team	(2) Attendee	(3) Host
Reported annual deaths (per million)			
Influenza	41 [80]	40 [78]	37 [68]
Cancer	12,349 [1,415]	12,300 [1,421]	11,620 [1,271]
Diabetes	4,218 [1,066]	4,171 [1,064]	3,879 [1,128]
Heart disease	29,132 [6,109]	29,104 [6,201]	27,210 [5,238]
Accidents	1,598 [1,136]	1,583 [1,126]	1,438 [1,034]
Suicide	305 [186]	304 [186]	346 [187]
Percentage age 65 or older	12 [4]	12 [4]	13 [5]
Percentage nonwhite	21 [14]	21 [13]	20 [12]
Mean daily temperature (degrees Fahrenheit)	58 [9]	58 [8]	66 [9]
Mean daily humidity (g/kg)	8 [2]	8 [2]	9 [3]
Mean daily precipitation (1/100 inches)	11 [4]	11 [4]	10 [6]
Number of counties	278	230	82
2009 population Age 65 +	15,618,641	13,710,357	5,776,926

Notes: Column 1 shows means among counties in cities that have ever hosted an NFL team. Column 2 shows means among counties in cities that have ever had teams that played in the Super Bowl. Column 3 shows means among counties in cities that have ever hosted the Super Bowl. Standard deviations are in brackets.

Table 1 shows means of outcome and control variables for various levels of involvement with the Super Bowl. We limit analysis to counties within an MSA with an NFL team. To obtain a balanced panel, we include all counties that ever had an NFL team during our sample period, even if the NFL removed or added the team at a later date: this includes Los Angeles (which had a team until 1994) and several expansion cities, most recently Nashville, which acquired a team in 1998. The resulting 278 counties span 30 states and cover approximately 50 percent of the US population. We also explore panels limited to cities with teams that ever attended the Super Bowl or cities that had an NFL team throughout the sample period.

TABLE 2. Effect of hosting the Super Bowl on local influenza mortality

	(1) Basic	(2) City-specific linear time trends	(3) City-specific quadratic time trends	(4) Limited to ever hosting	(5) Limited to continuous eligibility
Panel A: All ages					
Host	-0.616 (0.575)	-0.694 (0.652)	-0.699 (0.602)	-0.774 (0.599)	0.474 (0.661)
Mean	5.6	5.6	5.6	5.2	5.3
Percentage impact	-11	-12	-13	-15	9
Obs.	10,008	10,008	10,008	2,952	6,876
Panel B: Age 65 +					
Host	-3.861 (5.165)	-5.859 (6.163)	-5.262 (5.813)	-4.379 (5.344)	1.432 (6.416)
Mean	40.7	40.7	40.7	37.2	38.4
Percentage impact	-9	-14	-13	-12	4
Obs.	10,008	10,008	10,008	2,952	6,876

Notes: Coefficients indicate the effect of hosting the Super Bowl on local influenza deaths per million population. Regressions are at the county-season level, weighted by the appropriate population, and clustered by MSA. We assign counties as “hosting” by the county falling in the same MSA as the hosting stadium. Each regression includes controls for percentage nonwhite, percentage over age 65, and controls for the average number of days environment factors fell in these ranges: temperature in bins of 10 degrees Fahrenheit, humidity in bins of 2 g/kg, and precipitation in bins of 0.5 inches. All regressions include county and influenza season fixed effects. Standard errors shown in parentheses. ^a $p < 0.01$, ^b $p < 0.05$, ^c $p < 0.10$.

IV. Results

We use our core differences-in-differences identification strategy (equation 1) throughout. We first examine the impact of *hosting* the Super Bowl in Table 2. Being a host city is theoretically orthogonal to local team success, as the NFL determines Super Bowl location prior to the beginning of the relevant season. We find point estimates are negative but sufficiently noisy that we are unable to reject either negative or positive effects. We take this as suggestive evidence that Super Bowl–related travel does not pose a significant transmission risk for hosting cities. A number of effects might explain the negative coefficient: host cities may see an influx of income, raising health and health behavior, local mixing within hosting cities may not increase as a consequence of the Super Bowl game, or locals may actually decrease mixing in response to increased travelers (avoiding traffic, busy restaurants, etc.). These results suggest there is no net negative health effect of hosting a Super Bowl.

We next address the role of local team success, as measured by a local team *attending* the Super Bowl. Table 3 presents our findings. Panel A displays results across all ages and

TABLE 3. Effect of an affiliated team qualifying for the Super Bowl on local influenza mortality

	(1) Basic	(2) City-specific linear time trends	(3) City-specific quadratic time trends	(4) Limited to ever attending	(5) Limited to continuous eligibility
Panel A: All ages					
Attend	1.015 ^b (0.378)	0.915 ^b (0.403)	0.831 ^b (0.363)	1.033 ^b (0.379)	1.170 ^c (0.371)
Mean	5.6	5.6	5.6	5.6	5.3
Percentage impact	18	16	15	19	22
Obs.	10,008	10,008	10,008	8,280	6,876
Panel B: Age 65 +					
Attend	7.270 ^b (3.048)	5.977 ^a (3.232)	6.250 ^b (3.005)	7.368 ^b (3.028)	8.814 ^c (3.051)
Mean	40.7	40.7	40.7	40.1	38.4
Percentage impact	18	15	15	18	23
Obs.	10,008	10,008	10,008	8,280	6,876

Notes: Coefficients indicate the effect of sending an affiliated team to the Super Bowl on local influenza deaths per million population. Regressions are at the county-season level, weighted by the appropriate population, and clustered by MSA. We assign teams as “local” as measured by the county falling in the same MSA as the participating team. Each regression includes controls for percentage nonwhite, percentage over age 65, and controls for the average number of days environment factors fell in these ranges: temperature in bins of 10 degrees Fahrenheit, humidity in bins of 2 g/kg, and precipitation in bins of 0.5 inches. All regressions include county and influenza season fixed effects. Standard errors shown in parentheses. ^a $p < 0.01$, ^b $p < 0.05$, ^c $p < 0.10$.

panel B limits the sample to those age 65 and over. We find a local team making it to the Super Bowl increases influenza mortality in the participating county. The coefficient on *Attend* shows sending a team to the Super Bowl has a statistically significant effect for both all-age mortality and mortality among those age 65 and over. Focusing on the elderly, panel B indicates sending a team to the Super Bowl leads to an additional seven reported influenza deaths per million for those aged 65 and older in the home county. This is an 18 percent increase over the observed average (40.7 reported influenza deaths per million population age 65 and over). Point estimates are similar after adding city-specific linear (column 2) and quadratic (column 3) time trends. Column 4 limits the sample to counties with teams that ever attended the Bowl, which eliminates some counties associated with historically less-successful teams and newer expansion teams. In the more restricted sample our estimated impact is similar in precision and magnitude to that in the full sample. In all specifications other than column 2 of panel B, results are statistically

TABLE 4. Effect of an affiliated team qualifying for the Super Bowl on local influenza mortality over age 65, by dominant seasonal subtype or time between Super Bowl and influenza peak

	(1) Less than 24 days from peak	(2) 24 or more days from peak	(3) A(H3N2) or A(H2N2) dominated	(4) A(H1N1) or B dominated
Attend	11.087 ^c (3.082)	1.924 (5.958)	10.590 ^b (4.487)	1.972 (3.741)
Mean	38.3	43.3	53.1	22.6
Percentage impact	29	4	20	9
Obs.	5,282	4,726	5,838	4,170

Notes: Coefficients indicate the effect of sending an affiliated team to the Super Bowl on local influenza deaths per million population for those age 65 and over when the Super Bowl was held the given number of days from the influenza peak or during years when the dominant influenza strain was the type indicated. Regressions are at the county-season level, weighted by the appropriate population, and clustered by MSA. We assign teams as “local” as measured by the county falling in the same MSA as the participating team. Each regression includes controls for percentage nonwhite, percentage over age 65, and controls for the average number of days environment factors fell in these ranges: temperature in bins of 10 degrees Fahrenheit, humidity in bins of 2 g/kg, and precipitation in bins of 0.5 inches. All regressions include county and influenza season fixed effects. Standard errors shown in parentheses. ^a $p < 0.01$, ^b $p < 0.05$, ^c $p < 0.10$.

different from zero at a minimum of 5 percent, and statistically different from zero at a minimum of 10 percent in all specifications.

Table 4 restricts the sample to influenza seasons where we are either more or less likely to detect an effect. Column 1 breaks the sample into seasons where the Super Bowl occurred within 24 days of the period of highest annual influenza mortality nationwide.⁸ Column 2 limits the sample to seasons where the Super Bowl occurred 24 or more days away from the seasonal influenza mortality peak. In seasons where the Super Bowl was closer to the peak, mortality effects are over seven times larger (29 percent impact versus 4 percent impact). We next limit the sample to years where the dominant influenza subtype was especially virulent (column 3) or especially mild (column 4). The percentage effect of attending the Super Bowl on influenza mortality during seasons dominated by the more virulent A(H3N2) or A(H2N2) viruses (Greene, Ionides, and Wilson 2006) is approximately twice as large as the impact of attending the Super Bowl in a year where the dominant subtypes were the less virulent A(H1N1) or B viruses (Greene, Ionides, and Wilson 2006). Taken as a whole, the results in Table 4 support influenza as the vector of increased mortality.

8 We measure the peak of influenza mortality using the midpoint of the seven-day period with the highest influenza season mortality rate. We use 24 days as it comes closest to splitting our sample in half.

TABLE 5. Effect of an affiliated team qualifying for the Super Bowl on local influenza mortality over age 65, by risk or ICD

	(1) High-risk sending cities	(2) High-risk hosting cities	(3) ICD-8 (1974–78)	(4) ICD-9 (1979–98)	(5) ICD-10 (1999–2009)
Attend	14.477 ^b (6.089)	9.590 ^a (5.212)	13.226 (21.358)	5.166 (3.766)	3.070 (4.352)
Mean	61.8	49.4	129.4	36.9	19.6
Percentage impact	23	19	10	14	16
Obs.	5,004	5,282	1,390	5,560	3,058

Notes: Coefficients indicate the effect of sending an affiliated team to the Super Bowl on local influenza deaths per million population for those over age 65. Columns 1 and 2 stratify the sample by years when either the aggregate risk in the attending cities was in the top half over all years or years when the hosting cities were in the top half of all hosting cities. Columns 3, 4, and 5 stratify the sample by coding regime (ICD-8, ICD-9, or ICD-10) in mortality data. Regressions are at the county-season level, weighted by the appropriate population, and clustered by MSA. We assign teams as “local” as measured by the county falling in the same MSA as the participating team. Each regression includes controls for percentage nonwhite, percentage over age 65, and controls for the average number of days environment factors fell in these ranges: temperature in bins of 10 degrees Fahrenheit, humidity in bins of 2 g/kg, and precipitation in bins of 0.5 inches. All regressions include county and influenza season fixed effects. Standard errors shown in parentheses. ^a $p < 0.01$, ^b $p < 0.05$, ^c $p < 0.10$.

We next consider how effects vary by expected baseline risk and ICD classification. Columns 1 and 2 of Table 5 stratify by season-specific influenza risk. For each season, we calculate the all-age influenza death rate in the hosting MSA and average all-age death rate across both attending MSAs. Column 1 limits the sample to seasons when the population-weighted average influenza risk among all MSAs *sending* teams was in the top half of all rates. Column 2 limits the sample to years when the influenza death risk for MSAs *hosting* the game was in the top half of all rates. Team-sending cities experiencing contemporaneously high influenza mortality may be more likely to (1) pass influenza to other travelers at the Super Bowl game and (2) pass influenza locally at increased gatherings. The point estimate for influenza mortality in team-sending cities is slightly higher (23 percent versus 18 percent) during seasons when team-sending cities experience heightened mortality. Column 2 repeats this stratification for the hosting MSA. When hosting cities experience heightened influenza mortality, conditions may be more amenable to transmission of influenza from locals to travelers, and travelers may be more likely to return to their home cities infected. We find when cities above the median reported annual influenza mortality rate host the Super Bowl, the relationship between sending a Super Bowl team and local influenza mortality is effectively unchanged for sending cities as compared to the unrestricted sample (19 percent versus 18 percent). The effect of sending a team to the Super

Bowl is affected more by unusually high mortality rates in the area that sends the team rather than in the area the hosts the game.

Since our analysis spans several ICD schemes, columns 3, 4, and 5 of Table 5 restrict the sample to each regime, respectively. Results are consistent with previous estimates but imprecisely estimated. Point estimates are all positive and we cannot rule out impacts the same size as our main effect in each regime. In the restricted sample we cannot rule out the null hypothesis of no impact at conventional confidence levels.⁹ In stratifying by ICD regimes, we allow all included covariates to have differential effects by regime as well. The reader should keep this in mind when attempting comparison of effects across periods and with the full-sample estimate.

Thus far, our model bundles the impact of several aspects of successful team performance into a single estimate of sending a team to the Super Bowl. In Table 6, we examine the effects of the local team participating in the other postseason playoff games leading up to the Super Bowl. We note any playoff effects we detect are effectively attenuated Super Bowl effects. Mechanically, cities that send a team to the Super Bowl in a given season must also send a team to the playoffs in that season. Thus, our previously estimated effects include the joint impact of (1) sending a local team to the playoffs, (2) having a local team win the playoffs, and (3) sending a local team to the Super Bowl. The coefficient on *Attend* in our models measuring the impact of the playoffs includes (1) for all treated cities, and (2) and (3) for only two cities each year. Thus we interpret playoff effects as *part* of our prior estimates rather than *in addition to* said estimates. Unfortunately, we lack the power to simultaneously estimate both playoff and Super Bowl effects.

Playoffs are different than the Super Bowl in that one attending team's city is also the host. Whereas the Super Bowl is mostly hosted by warmer cities less amenable to influenza transmission, the higher-ranked playoff team hosts the playoff in their location. Our estimates of the impact of hosting any playoff games are the joint effect of (1) a successful local team and (2) effects of hosting a large event, including increased travel from outside or increased mixing by local residents. Table 6 shows that having a local team make even the playoffs correlates with an increase in influenza mortality. The effect is economically and statistically significant for hosting playoffs: having a local team host (and, by definition, compete in) a playoff game correlates with a 13 percent increase in local influenza mortality in our base specification, approximately three-fourths the size of our Super Bowl effect. Having a local team attend *but not host* a playoff game, however, has an economically smaller effect, around 5 percent and never achieving statistical significance at conventional levels. It is important to note the hosting playoff team is most often the team that also attends the Super Bowl. Playoff games were won by the hosting team 68 percent of the time in our sample, so part of the difference in stratification effects include playoff losers versus winners (and eventual Super Bowl participants).

9 Note the higher baseline recorded rate of influenza mortality (129 per million) in the ICD-8 regime will be differenced out by the time fixed effects. While the point estimates are larger during coding regimes where influenza is more frequently coded as the cause of death, the percentage impacts of attending the Super Bowl are similar across regimes.

TABLE 6. Effect of an affiliated team qualifying for the playoffs on local influenza mortality over age 65

	(1) Basic	(2) City-specific linear time trends	(3) City-specific quadratic time trends	(4) Limited to always eligible
Panel A: Attending any appearance				
	2.090 (1.967)	0.608 (1.678)	1.065 (1.551)	1.184 (1.889)
Mean	40.7	40.7	40.7	38.4
Percentage impact	5	2	3	3
Obs.	10,008	10,008	10,008	6,876
Panel B: Hosting higher rank				
	4.894 ^c (1.740)	3.819 (2.259)	3.967 ^a (1.988)	5.226 ^b (2.039)
Mean	39	39	39	36.8
Percentage impact	13	10	10	14
Obs.	10,008	10,008	10,008	6,876

Notes: Coefficients indicate the effect of either sending an affiliated team (for the lower-ranked team) to a Super Bowl playoff game or hosting an affiliated team (for the higher-ranked team) in a Super Bowl playoff game in a given season on influenza deaths per million population for those age 65 and over. Regressions are at the county-season level, weighted by the appropriate population, and clustered by MSA. We assign teams as “local” as measured by the county falling in the same MSA as the participating team. Each regression includes controls for percentage nonwhite, percentage over age 65, and controls for the average number of days environment factors fell in these ranges: temperature in bins of 10 degrees Fahrenheit, humidity in bins of 2 g/kg, and precipitation in bins of 0.5 inches. All regressions include county and influenza season fixed effects. Standard errors shown in parentheses. ^a $p < 0.01$, ^b $p < 0.05$, ^c $p < 0.10$.

Finally, Tables 7 and 8 present a number of placebo checks. In Table 7, we examine the effect of attending the Super Bowl on mortality either one or two seasons before or after the team actually attended. This tests for other, omitted, factors that influence trends in both playoff performance and influenza mortality. A correctly specified model should find no significant results. Accordingly, results are never significant at conventional levels, and point estimates are always lower than our main results.¹⁰ Table 8 presents results for various other mortality outcomes. With the exception of column 1, which includes all causes of death, these are noncommunicable diseases that should not necessarily be

¹⁰ Given standard errors, we can only rule out effects as large as the effect in column 1 of Table 3 at the 5 percent significance level for the effects two seasons earlier. This is a very stringent placebo check as a team’s success is serially correlated across seasons, and this season’s success affects subsequent years’ influenza mortality risk via increased playoff games.

TABLE 7. Placebo check: Effect of an affiliated team qualifying for the Super Bowl on local mortality in surrounding years

	(1) Two years earlier	(2) One year earlier	(3) One year later	(4) Two years later
Attend	-4.438 (4.474)	4.422 (4.164)	3.334 (2.951)	0.180 (4.559)
Mean	40.7	40.7	40.7	40.7
Percentage impact	-11	11	8	0
Obs.	10,008	10,008	10,008	10,008

Notes: Coefficients indicate the effect of having sent an affiliated team to the Super Bowl in the past (columns 1 and 2) or eventually sending a team in the future (columns 3 and 4) on local influenza deaths per million population for those age 65 and over. Regressions are at the county-season level, weighted by the appropriate population, and clustered by MSA. We assign teams as “local” as measured by the county falling in the same MSA as the participating team. Each regression includes controls for percentage nonwhite, percentage over age 65, and controls for the average number of days environment factors fell in these ranges: temperature in bins of 10 degrees Fahrenheit, humidity in bins of 2 g/kg, and precipitation in bins of 0.5 inches. All regressions include county and influenza season fixed effects. Standard errors shown in parentheses.

TABLE 8. Placebo check: Effect of an affiliated team qualifying for the Super Bowl on other local mortality causes

	(1) All	(2) Cancer	(3) Diabetes	(4) Suicide	(5) Heart	(6) Accident
Attend	133.623 (94.144)	-2.194 (23.109)	0.363 (26.013)	1.873 (4.826)	145.067 ^a (82.502)	23.769 (18.277)
Mean	49,962.1	12,406.7	4,039.4	322.1	29,506.6	1,646.2
Percentage impact	0	0	0	1	0	2
Obs.	10,008	10,008	10,008	10,008	10,008	10,008

Notes: Coefficients indicate the effect of sending an affiliated team to the Super Bowl on local deaths per million population for those age 65 and over by mortality cause. Regressions are at the county-season level, weighted by the appropriate population, and clustered by MSA. We assign teams as “local” as measured by the county falling in the same MSA as the participating team. Each regression includes controls for percentage nonwhite, percentage over age 65, and controls for the average number of days environment factors fell in these ranges: temperature in bins of 10 degrees Fahrenheit, humidity in bins of 2 g/kg, and precipitation in bins of 0.5 inches. All regressions include county and influenza season fixed effects. Standard errors shown in parentheses. ^a $p < 0.01$, ^b $p < 0.05$, ^c $p < 0.10$.

affected by infectious disease transmission. We rule out effects as large as the main effect on influenza in percentage terms for all investigated causes. We do observe slightly higher cardiovascular mortality, consistent with Warren-Gash, Smeeth, and Hayward (2009), who find suggestive evidence influenza increases cardiovascular-associated mortality.

V. Discussion and Conclusions

We find large, robust impacts of team Super Bowl participation on local influenza mortality, with suggestive effects for participation in playoff games as well. We hypothesize several plausible behavioral response explanations for our results. First, both playoffs and the Super Bowl increase travel to and from relevant cities. There may be increased transmission from fans that travel to the games themselves or, in the case of playoff games (hosted by one of the participating teams), that are exposed to incoming travelers. Fans mix with other travelers at the airport, and with fans of the opposing team and each other while at the games. Without detailed data on travel patterns we are unable to directly isolate the travel mechanism, though we find no impacts of *hosting* the Super Bowl on local influenza mortality, which suggests travelers from participating cities are not infecting the local hosting population. Although the selection process is secretive and owners' preferences are unknown, weather is likely an important factor in location choice (Pedulla 2012). Figure 2 shows the game has been most frequently hosted by cities with relatively warmer, more humid climates, whereas attending teams come from all weather regions, as Figure 1 shows. Such warmer, humid climates serve as barriers to influenza transmission and thus mute the effect of the attendees on local mortality (Lowen et al. 2007; Shaman and Kohn 2009; Barreca and Shimshack 2012). When we limit our analysis of the effects of *attending* the Super Bowl to generally warmer cities that have *hosted* the Super Bowl we see little response, which further supports this explanation. Travelers may not mix intimately enough with the local population, but return and interact closely with relatives in vulnerable populations. Hosting cities may increase public health activities in anticipation of the Super Bowl (Lombardo et al. 2008), which may mitigate transmission, whereas attending cities may not scale up public health efforts since the gathering is not held in their cities. These activities may raise media awareness of risk in hosting cities and cause individuals to engage in voluntary defensive behavior (Fenichel, Kuminoff, and Chowell 2013), but raise no media awareness and defensive practices among inhabitants from participating cities.¹¹ We note any effects we detect are net of avoidance behavior on behalf of those in treatment cities. While socialization among the general population might increase as a result of team success, individuals may actively limit their contact with sensitive populations in an attempt to decrease disease vectors, or decrease travel in an attempt to avoid

11 Fans of attending cities may be subjected to additional physiological stresses that make each individual more susceptible to influenza infection (Kiecolt-Glaser et al. 1996). We detected slightly elevated cardiovascular mortality in attending cities, which is consistent with a stress-related mechanism (Kivimäki et al. 2002). Risky behavior may also be at play, as evidenced by increased rates of low birth weight and greater alcohol and tobacco consumption in areas with successful Super Bowl teams (Duncan, Mansour, and Rees 2013).

increased contact with other travelers. Assuming avoidance behavior is more drastic in treatment counties, this will push our estimates toward zero. If individuals are already engaging in the lower-cost avoidance behavior, this would decrease the effectiveness of any further policy that is avoidance behavior-based.

As an alternate possibility, increased success of an NFL team is cause for increased gathering and celebration by local inhabitants. Friends and family are more likely to gather to watch games either at public venues such as bars and restaurants or at private venues such as “Super Bowl parties,” sharing food and beverages and generally increasing contact. Such person-to-person contact can shift influenza transmission prior to any actual Super Bowl–related travel. Our findings support this mechanism, as we find suggestive evidence of increased mortality even among counties associated with teams that make it to playoff games, but not the Super Bowl. We do find larger (and more precisely estimated) impacts for teams that are higher ranked for at least one playoff game, again suggesting general team success (and associated increased local gatherings) is an important vector of transmission. These higher-ranked teams are also more likely to make the Super Bowl, and may have a more active fan base than counties with teams that merely attend the playoffs.¹² Direct comparison of playoff game impacts and Super Bowl impacts are difficult, as playoff games are necessarily earlier in the year when influenza transmission probability is generally lower. In this sense, smaller estimated effects for playoff success are consistent with our estimates that indicate the Super Bowl has less impact on mortality in counties associated with teams that participate in the Super Bowl when the bowl is held further from the peak of influenza season, which may point to welfare gains from moving the date of the Super Bowl away from the peak of high mortality influenza seasons.

In sum, we present evidence influenza mortality increases in cities with NFL teams during successful postseason play. Our work suggests having a successful NFL team can impose a negative externality on the MSA’s elderly population, potentially through increased social gatherings, increased local travel, or both. Such effects extend to other similar events, such as the Olympics, World Cup (soccer), and World Series (baseball). Strategies exist to help offset such effects. If a major contributor to increased influenza spread is local gatherings for watching games, a simple policy solution is to increase awareness of influenza transmission vectors during times of sports-related gatherings. Reminding people to wash their hands and avoid sharing drinks or food at parties during the height of influenza season, especially if they have high amounts of contact with vulnerable populations, could have large social returns.

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12 We also find slightly larger, though not statistically different, impacts for Super Bowl winners than Super Bowl losers, providing further suggestive evidence it is the successful teams driving these effects. These additional results are omitted for brevity and available upon request.

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